The Long-Run Phillips Curve is... a Curve ¹

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Inflation: Drivers and Dynamics Conference 2021

Cleveland Fed and ECB

 $^{^1 {\}rm Views}$ expressed are those of the authors and do not necessarily reflect official positions of De Nederlandsche Bank

The question

An old debate: is there any trade-off between inflation and output/unemployment in the long run?

Phelps (1967), Friedman (1968): Natural rate hypothesis

"there is no permanent trade-off":
=>the long-run Phillips curve is vertical

- Cornerstone role in macroeconomic theory and practice
- The working assumption of central banks in the implementation of monetary policy

The question

It is surprising to note that:

Empirically: There is little econometric work devoted to test the absence of a long-run trade-off.

Some literature: King and Watson (1994); Beyer and Farmer (2007); Berentsen et al. (2011); Haug and King (2014); Benati (2015)

- Theoretically: Modern macroeconomic sticky price frameworks generally do not imply the absence of a long- run relation
 - The Generalized NK model delivers a negative relationship between steady state inflation and output. See Ascari (2004); Ascari and Sbordone (2014)

Results

What is the long-run relation between inflation and output?

1. Time series model

- The LRPC is not vertical, it is negatively sloped (higher inflation is related to lower output in the LR)
- The key to get this result: model the LRPC as non linear
- Methodological contribution: a "convenient" non-linear approach
- 2. Structural model
 - GNK model (Ascari and Ropele, 2009; Ascari and Sbordone, 2014): higher trend inflation causes lower GDP in the LR
 - The model has the two key features from the statistical analysis: non-linear and negatively sloped LRPC
 - The model is also able to capture the quantitative features of the time series analysis

The time series approach: A time-varying equilibrium VAR

Generalization of Steady State VAR (Villani, 2009; Del Negro et al., 2017; Johannes and Mertens, 2021):

$$A(L) (X_t - \bar{X}_t) = \varepsilon_t \quad \varepsilon_t \sim N(0, \Sigma_{\varepsilon, t})$$
(1)

X_t is a (n × 1) vector with observed variables at time
 X
 x x t is the vector with the long-run values of X_t

Trend-cycle decomposition:

$$X_t = \bar{X}_t + \hat{X}_t$$



The model

- Three observables: GDP per capita, inflation and interest rate
- The short-run component: VAR with 4 lags

THE MODEL FOR THE LONG RUN

$$\begin{split} \bar{y}_t &= y_t^* + \delta(\bar{\pi}_t) \quad \text{the equilibrium level of output as function of inflation} \\ y_t^* &= y_{t-1}^* + g_t + \eta_t^y \\ g_t &= g_{t-1} + \eta_t^g \\ \delta(\bar{\pi}_t) : \delta(0) &= 0 \end{split}$$

 $ar{\pi}_t = ar{\pi}_{t-1} + \eta_t^\pi$ trend inflation is random walk

 $ar{i}_t = ar{\pi}_t + cg_t + z_t$ long-run Fisher equation $z_t = z_{t-1} + \eta_t^z$

A non-linear long-run Phillips curve

Our choice of $\delta(\bar{\pi}_t)$ is a piecewise linear function:

$$\begin{split} \bar{y}_t &= y_t^* + \delta(\bar{\pi}_t) \\ \delta(\bar{\pi}_t) &= \begin{cases} k_1 \bar{\pi}_t & \text{if } \bar{\pi}_t \leq \tau \\ k_2 \bar{\pi}_t + c_k & \text{if } \bar{\pi}_t > \tau \end{cases} \end{split}$$

- It is simpler to treat: methodological contribution
- It can approximate the kind of non-linearity we have in mind without imposing strong assumptions on a specific functional form
- It is easy to interpret

A piecewise linear approach

The model can be written in state space form:

$$Y_{t} = D(\theta_{t}) + F(\theta_{t})\theta_{t} + \epsilon_{t}$$

$$\theta_{t} = M(\theta_{t}) + G(\theta_{t})\theta_{t-1} + P(\theta_{t})\eta_{t}$$
(2)
(3)

where, in particular

$$(D, F, M, G, P) = \begin{cases} (D_1, F_1, M_1, G_1, P_1) & \text{if } \bar{\pi}_t \le \tau \\ (D_2, F_2, M_2, G_2, P_2) & \text{if } \bar{\pi}_t > \tau \end{cases}$$
(4)

- Methodological contribution: we find the likelihood and the posterior distribution of θ_t analytically
- Compromise between efficiency and misspecification

Estimation

- US data, sample from 1960Q1 to 2008Q2
- Bayesian approach

Two sources of non linearity: stochastic volatility and a piecewise linear LRPC => Particle filtering approach

- 1. "Rao-Blackwellization", thanks to the analytical results on the piecewise linear model
- 2. Particle filtering also to approximate the posterior distribution of the parameters
 - Particle learning by Carvalho Johannes Lopes and Polson (2010); see also Mertens and Nason (2020)
 - Mixture of Normal distributions as approximation of the posterior of τ (Liu and West, 2001)

Strategy for parameter learning follows Chen, Petralia and Lopes (2010) and Ascari, Bonomolo and Lopes (2019)

Estimation results - Linear model

A vertical (or flat) long-run Phillips curve



Figure: Posterior distributions of the slope of the LRPC - Linear model.

Estimation results - Non-linear model

Non linear and negatively sloped long-run Phillips curve



Figure: Posterior distributions of the slopes of the LRPC - Non-linear model.

Estimation results - Non-linear model

The threshold:



Figure: Posterior distributions of τ - Non-linear model.

Estimation results - non-linear model

A non-linear, negatively sloped long-run Phillips curve



Figure: LRPC - Non-linear model. Median and 90% probability interval.

Estimation results - non-linear model



Figure: Inflation and trend inflation - Non linear model.

The cost of trend inflation: the long-run output gap

$$\hat{Y}_t = \frac{Y_t}{\bar{Y}_t} = \frac{Y_t}{Y_t^*} \frac{Y_t^*}{\bar{Y}_t}$$
(5)



Figure: Long-run output gap estimated through the non-linear model.

The structural model

- A variant of Ascari and Ropele (2009), Ascari and Sbordone (2014) GNK model:
 - Inter-temporal Euler equation featuring (external) habit formation in consumption
 - Generalized New Keynesian Phillips curve featuring positive trend inflation
 - Taylor-type monetary policy rule
- Time varying trend inflation => LRPC is:
 - Non-linear
 - Negatively sloped
- When taking decisions the agents consider trend inflation as a constant parameter: anticipated-utility model (Kreps, 1998; Cogley and Sbordone, 2008)
- Stochastic volatility to the four shocks: discount factor, technology, monetary policy and trend inflation

The costs of trend inflation

- Price stickiness => price dispersion and inefficiency in the quantity produced
- Higher trend inflation leads to higher price dispersion and increases output inefficiency

Formally:

$$N_t = \int_0^1 N_{i,t} di = \int_0^1 \left(\frac{Y_{i,t}}{A_t}\right)^{\frac{1}{1-\alpha}} di = \underbrace{\int_0^1 \left(\frac{P_{i,t}}{P_t}\right)^{\frac{-\varepsilon}{1-\alpha}}}_{s_t} di \left(\frac{Y_t}{A_t}\right)^{\frac{1}{1-\alpha}}$$

Aggregate output is:

$$Y_t = \frac{A_t}{s_t^{1-\alpha}} N_t^{1-\alpha}$$

with long-run price dispersion: $\bar{s}_t = g(\bar{\pi}_t)$



Comparing long-run Phillips curves: VAR and GNK

The GNK model measures the costs of trend inflation consistently with the VAR $% \left({{\rm{AR}}} \right)$



Figure: Long-run Phillips curve: median (continuous line) and 90% probability interval (dashed lines) - comparison between VAR (blue) and GNK (black) estimates.

Conclusions

- What is the long-run relation between inflation and output?
- A time series model suggests that the LRPC is:
 - Non linear
 - Negatively sloped
- ▶ We interpret these findings through the lens of a GNK model
- This model is able to measure the costs implied by the LRPC consistently with the time series model

EXTRA

Econometric strategy

We use a particle filtering strategy to approximate the joint posterior distribution of latent processes and parameters:

Latent processes: a "conditional piecewise linear model"

$$p\left(\theta_{t}, \Sigma_{\epsilon, t} | Y_{t}\right) = \underbrace{p\left(\theta_{t} | \Sigma_{\epsilon, t}, Y_{t}\right)}_{\text{"optimal importance kernel"}} \underbrace{p\left(\Sigma_{\epsilon, t} | Y_{t}\right)}_{\text{"blind proposal"}}$$

Parameters:

- Particle learning by Carvalho Johannes Lopes and Polson (2010); see also Mertens and Nason (2020)
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A fully adapted particle filter

At
$$t - 1$$
: $\left\{\theta_{t-1}^{(i)}\right\}_{1}^{N}$ approximate $p\left(\theta_{t-1} | \psi, X_{1:t-1}\right)$

1. Resample

$$\begin{aligned} & \bullet \quad \text{Compute } \tilde{w}_t^{(i)} \propto p\left(X_t | \theta_{t-1}^{(i)}, \psi, X_{1:t-1}\right) \\ & \bullet \quad \text{Resample } \left\{\tilde{\theta}_{t-1}^{(i)}\right\}_1^N \text{ using } \left\{\tilde{w}_t^{(i)}\right\}_1^N \end{aligned}$$

2. Propagate

• draw
$$\theta_t^{(i)} \sim p\left(\theta_t | \tilde{\theta}_{t-1}^{(i)}, \psi, X_{1:t-1}\right)$$

where:

p (X_t|θ⁽ⁱ⁾_{t-1}, ψ, X_{1:t-1}) is a weighted sum of Unified Skew Normal distributions (Arellano-Valle and Azzalini, 2006)
 p (θ_t|θ⁽ⁱ⁾_{t-1}, ψ, X_{1:t-1}) is a weighted sum of multivariate truncated Normal distributions



Household

The economy is populated by a representative agent with utility

$$E_0 \sum_{t=0}^{\infty} \beta^t d_t \left[\ln \left(C_t - h \mathbf{C}_{t-1} \right) - d_n \frac{N_t^{1+\varphi}}{1+\varphi} \right]$$

Budget constraint is given by

$$P_t C_t + R_t^{-1} B_t = W_t N_t + D_t + B_{t-1}$$

 d_t is a discount factor shock which follows an AR(1) process

$$\ln d_t = \rho_d \ln d_{t-1} + \epsilon_{d,t}$$



Final good firm

Perfectly competitive final good firms combine intermediate inputs

$$Y_t = \left[\int_0^1 Y_{i,t}^{\frac{\varepsilon-1}{\varepsilon}} di\right]^{\frac{\varepsilon}{\varepsilon-1}} \qquad \varepsilon > 1$$

Price index is a CES aggregate of intermediate input prices

$$P_t = \left[\int_0^1 P_{i,t}^{1-\varepsilon} di\right]^{\frac{1}{1-\varepsilon}}$$

The demand schedule for intermediate input

$$Y_{i,t} = \left[\frac{P_{i,t}}{P_t}\right]^{-\varepsilon} Y_t$$



Intermediate good firm

Each firm i produces according to the production function

$$Y_{i,t} = A_t N_{i,t}^{1-\alpha}$$

where A_t denotes the level of technology and its growth rate $g_t \equiv A_t/A_{t-1}$ follows

$$\ln g_t = \ln \overline{g} + \epsilon_{g,t}$$



Price setting

Firms adjust prices $P^*_{i,t}$ to maximize expected discounted profits with probability 0 $<1-\theta<1$

$$E_{t}\sum_{j=0}^{\infty}\theta^{j}\beta^{j}\frac{\lambda_{t+j}}{\lambda_{t}}\left[\frac{P_{i,t}^{*}}{P_{t+j}}Y_{i,t+j}-\frac{W_{t+j}}{P_{t+j}}\left[\frac{Y_{i,t+j}}{A_{t+j}}\right]^{\frac{1}{1-\alpha}}\right]$$

subject to the demand schedule

$$Y_{i,t+j} = \left[\frac{P_{i,t}^*}{P_{t+j}}\right]^{-\varepsilon} Y_{t+j},$$

where λ_t is the marginal utility of consumption.

Back

The Phillips curve

The first order condition for the optimized relative price $x_t (= \frac{P_{i,t}^*}{P_t})$ is given by

$$(x_t)^{1+\frac{\epsilon\alpha}{1-\alpha}} = \frac{\varepsilon}{(\varepsilon-1)(1-\alpha)} \frac{E_t \sum_{j=0}^{\infty} (\theta\beta)^j \lambda_{t+j} \frac{W_{t+j}}{P_{+j}} \left[\frac{Y_{t+j}}{A_{t+j}}\right]^{\frac{1}{1-\alpha}} \pi_{t|t+j}^{\frac{\varepsilon}{(1-\alpha)}}}{E_t \sum_{j=0}^{\infty} (\theta\beta)^j \lambda_{t+j} \pi_{t|t+j}^{\varepsilon-1} Y_{t+j}}$$

where
$$\pi_{t|t+j} = \frac{P_{t+1}}{P_t} \times ... \times \frac{P_{t+j}}{P_{t+j-1}}$$
 for $j \ge 1$ and $\pi_{t|t} = \pi_t$.

Back

Price setting contd.

Aggregate price level evolves according to

$$P_t = \left[\int_0^1 P_{i,t}^{1-\varepsilon} di\right]^{\frac{1}{1-\varepsilon}} \Rightarrow$$
$$x_t = \left[\frac{1-\theta\pi_t^{\varepsilon-1}}{1-\theta}\right]^{\frac{1}{1-\varepsilon}}.$$

Finally, price dispersion $s_t \equiv \int_0^1 (\frac{P_{i,t}}{P_t})^{-\varepsilon} di$ can be written recursively as:

$$s_t = (1 - \theta) x_t^{-\varepsilon} + \theta \pi_t^{\varepsilon} s_{t-1}$$



Monetary policy

$$\frac{R_t}{\overline{R}_t} = \left(\frac{R_{t-1}}{\overline{R}_t}\right)^{\rho} \left[\left(\frac{\pi_t}{\overline{\pi}_t}\right)^{\psi_{\pi}} \left(\frac{Y_t}{Y_t^n}\right)^{\psi_{\chi}} \left(\frac{g_t^y}{\overline{g}}\right)^{\psi_{\Delta y}} \right]^{1-\rho} e^{\varepsilon_{r,t}}$$

$$\ln \overline{\pi}_t = \ln \overline{\pi}_{t-1} + \epsilon_{\overline{\pi},t}$$

where $\overline{\pi}_t$ denotes trend inflation, Y_t^n is the flex-price output and g_t^y is growth rate of output.



Estimates of the parameters

Parameter		Prior		Posterior
	Density	Mean	St Dev	
ψ_{π}	Gamma	1.5	0.5	2.05 [1.83 2.3]
ψ_x	Gamma	0.125	0.05	$\begin{array}{c} 0.1 \\ [0.06 \ 0.16] \end{array}$
$\psi_{\Delta y}$	Gamma	0.125	0.05	$\begin{array}{c} 0.42 \\ \left[0.26 \ 0.67 ight] \end{array}$
ho	\mathbf{Beta}	0.7	0.1	$\begin{array}{c} 0.74 \\ \left[0.71 \ 0.77 ight] \end{array}$
h	Beta	0.5	0.1	$\underset{[0.36 0.46]}{0.41}$
r^*	Gamma	2	0.5	1.88 [1.68 2.1]
θ	Beta	0.5	0.1	0.51 [0.46 0.56]
$ ho_d$	Beta	0.7	0.1	$\begin{array}{c} 0.79 \\ \left[0.74 \ 0.83 ight] \end{array}$
	Density	Mean	Degree of freedom	
δ_d^2	Inverse Gamma	0.02^{2}	5	$\begin{array}{c} 0.047^2 \\ [0.042^2 \ 0.054^2] \end{array}$
δ_g^2	Inverse Gamma	0.02^{2}	5	$\begin{array}{c} 0.047^2 \\ 0.042^2 & 0.053^2 \end{array}$
δ_r^2	Inverse Gamma	0.02^{2}	5	$\begin{array}{c} 0.029^2 \\ 0.026^2 \ 0.033^2 \end{array}$
δ_π^2	Inverse Gamma	0.02^{2}	5	0.012^2 [0.011 ² 0.014 ²]

Table 1: Prior and Posterior Distributions

Posterior median and 90% credibility interval in brackets